

# Math 3280 Tutorial 7

Recall. 1. Find the distribution of  $g(X)$ ,  $X$  is a cts r.v. with pdf  $f_X$ ,  $g: \mathbb{R} \rightarrow \mathbb{R}$ .

$$F_{g(X)}(y) = P(g(X) \leq y)$$

$$\Downarrow f_{g(X)}(y) = \frac{dF_{g(X)}(y)}{dy}$$

2. joint cdf of  $X, Y$ ,  $F(a, b) := P(X \leq a, Y \leq b)$ ,  $a, b \in \mathbb{R}$ .

$$F_X(a) = \lim_{b \rightarrow \infty} F(a, b)$$

$$F_Y(b) = \lim_{a \rightarrow \infty} F(a, b)$$

$$P(X > a, Y > b) = 1 - F(a, \infty) - F(\infty, b) + F(a, b)$$

3. For discrete r.v.s  $X, Y$ , the probability mass function

$$p(x, y) = P(X=x, Y=y)$$

$$P_X(x) = \sum_{y \in \mathcal{Y}} p(x, y)$$

4. Two r.v.s are jointly continuous if exists

$$f: \mathbb{R}^2 \rightarrow [0, \infty)$$

s.t.

$$P((X, Y) \in C) = \iint_C f(x, y) dx dy$$

$$C \subseteq \mathbb{R}^2$$

$$P(X \leq x, Y \leq y) = \int_0^y \int_0^x f(x, y) dx dy$$

Example 1: If  $X$  is uniformly distributed on  $(0,1)$ . what is the probability density function of  $Y=e^X$ .

Solution:

$$f_X(x) = \begin{cases} 1 & x \in (0,1) \\ 0 & \text{o.w.} \end{cases}$$

$$Y = e^X \in (1, e).$$

$$\text{If } x \leq 1, F_Y(x) = P(Y \leq x) = 0.$$

$$\text{If } x \geq e, F_Y(x) = P(Y \leq x) = 1.$$

$$\text{If } \underline{1 < x < e}, F_Y(x) = P(Y \leq x) = P(e^X \leq x)$$

$$\underline{0 < \ln x < 1}$$

$$= P(X \leq \ln x)$$

$$= \int_0^{\ln x} 1 dt = \ln x$$

$$f_Y(x) = \begin{cases} 0 & x \leq 1 \\ \frac{1}{x} & x \in (1, e) \\ 0 & x \geq e. \end{cases}$$

Example 2: The median of a r.v. with distribution  $F$  is defined to be the value of  $m$  such that  $F(m) = \frac{1}{2}$ . Find the median of  $X$  if  $X$  is

- uniformly distributed on  $(a,b)$
- normal r.v. with  $\mu, \sigma^2$ .
- exponential r.v. with  $\lambda$ .

Solution: (a)  $X \sim U(a,b)$

$$f_X(x) = \begin{cases} \frac{1}{b-a} & x \in (a,b) \\ 0 & \text{o.w.} \end{cases}$$

$$F_X(m) = \frac{1}{2}, \quad m \in (a, b)$$

$$F_X(m) = \int_a^m f_X(x) dx = \frac{m-a}{b-a} = \frac{1}{2} \Rightarrow m = \frac{a+b}{2}$$

(a).  $X \sim N(\mu, \sigma^2)$ .  $\sigma > 0$ .

$$f_X(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

$$\frac{1}{2} = F(m) = \int_{-\infty}^m f_X(x) dx = P(X \leq m)$$

$$(Y = \frac{X-\mu}{\sigma}) = P(Y \leq \frac{m-\mu}{\sigma})$$

$$Y \sim N(0,1) = \Phi\left(\frac{m-\mu}{\sigma}\right)$$

cdf of  $Y$ .

$$\Phi(0) = \frac{1}{2}$$

$$\frac{m-\mu}{\sigma} = 0 \Rightarrow m = \mu$$

$$(c). f_X(x) = \begin{cases} \lambda e^{-\lambda x} & x \geq 0 \\ 0 & \text{o.w.} \end{cases}$$

$$\frac{1}{2} = F(m) = \int_0^m \lambda e^{-\lambda x} dx = 1 - e^{-\lambda m}$$

$$\Rightarrow \lambda m = \ln 2$$

$$m = \frac{\ln 2}{\lambda}$$

3. For any real number  $y$ , define  $y^+$  by

$$Y = \begin{cases} Y & Y \geq 0 \\ 0 & \text{o.w.} \end{cases}$$

Let  $c$  be a constant.

(a) Show that

$$E[(Z-c)^+] = \frac{1}{\sqrt{2\pi}} e^{-\frac{c^2}{2}} - c(1 - \Phi(c))$$

cdf of  $N(0,1)$

$Z$  is a standard normal r.v.

(b)  $Z \sim N(\mu, \sigma^2)$ , find  $E[(Z-c)^+]$ .

Solution: (a)  $Z \sim N(0,1)$

$$f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}$$

$$E[(Z-c)^+] = \int_c^\infty (z-c) f_Z(z) dz, \quad z < c, (z-c)^+ = 0.$$

$$= \int_c^\infty z \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz - c \int_c^\infty \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz$$

$$\int_c^\infty \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} d\left(\frac{z^2}{2}\right)$$

$$= \frac{1}{\sqrt{2\pi}} (e^{-\frac{z^2}{2}} / \infty)$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{c^2}{2}} / c$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{c^2}{2}} - c(1 - \Phi(c))$$

$$(b) f_Z(z) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(z-\mu)^2}{2\sigma^2}}, \quad (\sigma \geq 0)$$

$$E[(Z-c)^+] = \int_c^\infty (z-c) \cdot \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(z-\mu)^2}{2\sigma^2}} dz$$

$$\frac{y-\mu}{\sigma} \int_{\frac{c-\mu}{\sigma}}^{\infty} (y+\mu) \frac{1}{\sqrt{\sigma}} e^{-\frac{y^2}{2}} dy$$

$$= \sigma \int_{\frac{c-\mu}{\sigma}}^{\infty} \frac{1}{\sqrt{\sigma}} (y - \frac{c-\mu}{\sigma}) \cdot e^{-\frac{y^2}{2}} dy$$

denote  $\frac{c-\mu}{\sigma} = a$

$$= \sigma \cdot \int_a^{\infty} \frac{1}{\sqrt{\sigma}} (y-a) e^{-\frac{y^2}{2}} dy$$

$$= \sigma \cdot E[(Y-a)^+], \quad Y \sim N(0,1)$$

$$= \sigma \cdot \left( \frac{1}{\sqrt{2\pi}} e^{-\frac{a^2}{2}} - a \cdot (1 - \Phi(a)) \right).$$

Example 4. The joint density function of  $X$  and  $Y$  is given by

$$f(x,y) = \begin{cases} 2e^{-x} \cdot e^{-2y} & , 0 < x < \infty, 0 < y < \infty \\ 0 & , \text{o.w.} \end{cases}$$

(a) Find

$$P(X > 1, Y < 1)$$

(b)  $P(X < Y)$

(c)  $P(X < 1)$

$$(c) P(X < a)$$

$$\text{Solution: (a) } P(X > 1, Y < 1) = \int_0^1 \left( \int_1^{\infty} f(x, y) dx \right) dy$$

$$= \int_0^1 \left( \int_1^{\infty} 2e^{-x} \cdot e^{-2y} dx \right) dy$$

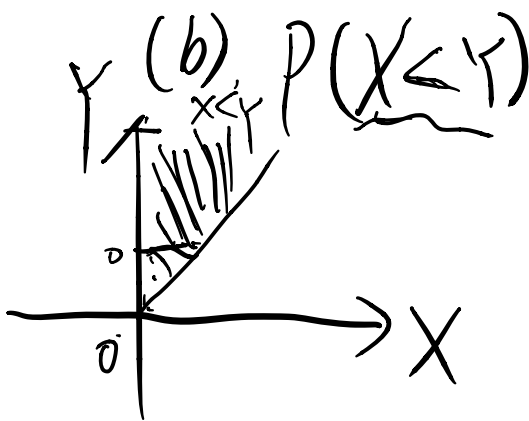
$$\int_1^{\infty} e^{-x} dx = e^{-1}$$

$$= \int_0^1 2e^{-2y} \cdot e^{-1} dy$$

$$= e^{-1} \int_0^1 2e^{-2y} dy$$

$$= e^{-1} \cdot \left( -e^{-2y} \Big|_0^1 \right)$$

$$= e^{-1} \cdot (1 - e^{-2})$$



$$P(X < Y) = \int_0^{\infty} \left( \int_0^y f(x, y) dx \right) dy$$

$$= \int_0^{\infty} \left( \int_0^y 2e^{-x} e^{-2y} dx \right) dy$$

$$2e^{-2y} \cdot \int_0^y e^{-x} dx = 2e^{-2y} \cdot (1 - e^{-y})$$

$$= \int_0^{\infty} 2e^{-2y} \cdot (1 - e^{-y}) dy$$

$$= 1 - \frac{2}{3} = \frac{1}{3}$$

$$(c). P(X < a) = \int_0^{\infty} \int_0^a f(x,y) dx dy$$

$$= \int_0^{\infty} \left( \int_0^a 2e^{-x} e^{-2y} dx \right) dy$$

$$= \int_0^{\infty} 2 \cdot e^{-2y} \cdot \underbrace{\left( \int_0^a e^{-x} dx \right)}_{1 - e^{-a}} dy$$

$$= \int_0^{\infty} 2 \cdot e^{-2y} (1 - e^{-a}) dy$$

$$= (1 - e^{-a}) \cdot \underbrace{\int_0^{\infty} 2e^{-2y} dy}_{\lambda=2, \text{ exponential}} = 1$$

$$= 1 - e^{-a}$$